QTL Mapping, MAS, and Genomic Selection

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A short-course organized by

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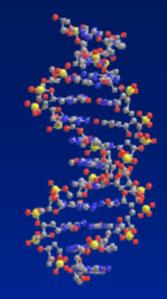


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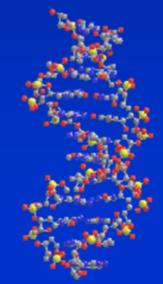






Linkage Disequilbrium to Genomic Selection







Course overview

- Day 1
 - Linkage disequilibrium in animal and plant genomes
- Day 2
 - QTL mapping with LD
- Day 3
 - Marker assisted selection using LD
- Day 4
 - Genomic selection
- Day 5
 - Genomic selection continued

Mapping QTL using LD

- Association testing with single marker regression
- Accounting for population structure
- LD mapping with haplotypes
- The Identical by descent (IBD) approach
- Combined linkage-linkage disequilibrium mapping

Mapping QTL with LD

- LD mapping of QTL exploits population level associations between markers and QTL.
 - Associations arise because there are small segments of chromosome in the current population which are descended from the same common ancestor
 - These chromosome segments, which trace back to the same common ancestor without intervening recombination, will carry identical marker alleles or marker haplotypes
 - If there is a QTL somewhere within the chromosome segment, they will also carry identical QTL alleles
- The simplest mapping strategy to exploit LD is a genome wide association test using single marker regression.

 Association between a marker and a trait can be tested with the model

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}} \boldsymbol{\mu} + \mathbf{X} g + \mathbf{e}$$

• Where

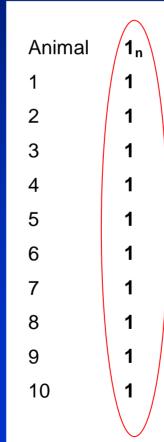
- y is a vector of phenotypes
- 1n is a vector of 1s allocating the mean to phenotype,
- X is a design matrix allocating records to the marker effect,
- g is the effect of the marker
- e is a vector of random deviates ~ $N(0,\sigma_e^2)$
- Underlying assumption here is that the marker will only affect the trait if it is in linkage disequilibrium with an unobserved QTL.

• A small example

Animal	Phenotpe	SNP allele 1	SNP allele
1	2.030502	1	1
2	3.542274	1	2
3	3.834241	1	2
4	4.871137	2	2
5	3.407128	1	2
6	2.335734	1	1
7	2.646192	1	1
8	3.762855	1	2
9	3.689349	1	2
10	3.685757	1	2

• The design vector **1**_n allocates phenotypes to the mean

Animal	Phenotpe	SNP allele 1	SNP allele
1	2.030502	1	1
2	3.542274	1	2
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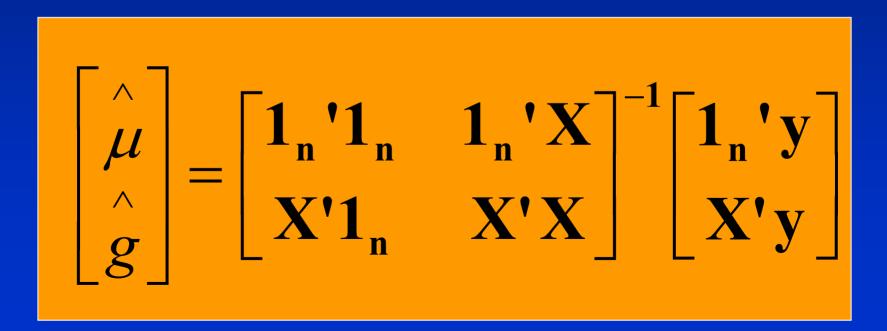
- The design vector $\mathbf{1}_{n}$ allocates phenotypes to the mean
- The design vector **X** allocates phenotypes to genotypes

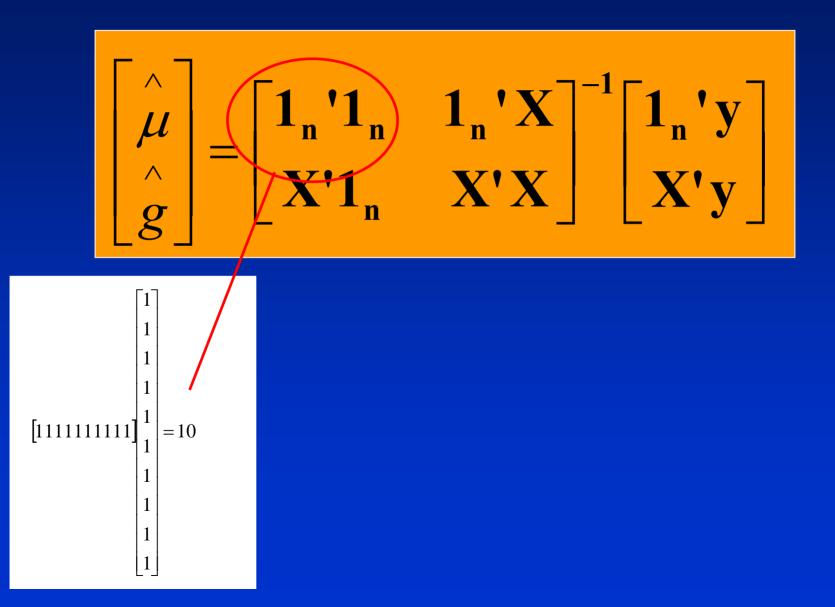
						X, Number of "2"
Animal	Phenotpe	SNP allele 1	SNP allele	Animal	1 _n	alleles
1	2.030502	1	1	1	1	0
2	3.542274	1	2	2	1	1
3	3.834241	1	2	3	1	1
4	4.871137	2	2	4	1	2
5	3.407128	1	2	5	1	1
6	2.335734	1	1	6	1	0
7	2.646192	1	1	7	1	0
8	3.762855	1	2	8	1	1
9	3.689349	1	2	9	1	1
10	3.685757	1	2	10	1	1

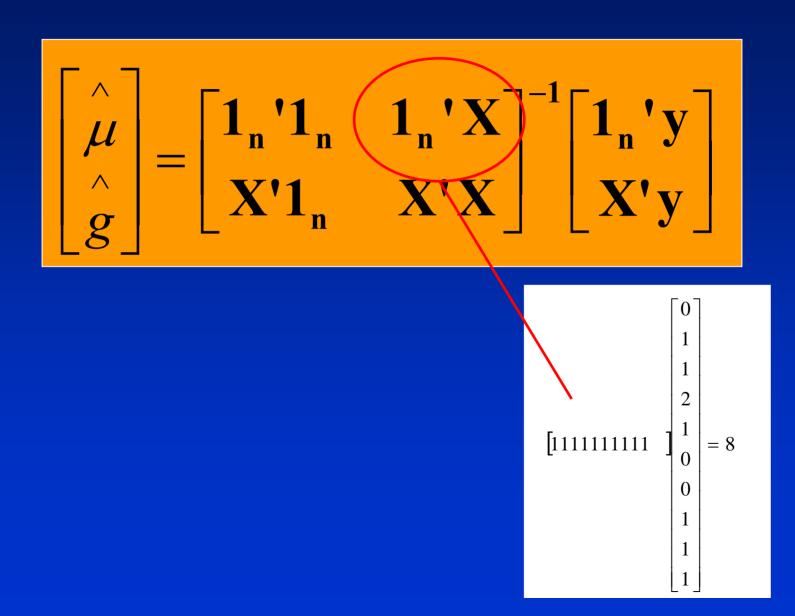
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					X, Number of "2"
Anim	al Phenotpe SNP allele 1	SNP allele	Animal	1 _n	alleles
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4	4.871137 2	2	4	1	2
5	3.407128 1	2	5	1	1
6	2.335734 1	1	6	1	0
7	2.646192 1	1	7	1	0
8	3.762855	2	8	1	1
9	3.689349 1	2	9	1	1
10	3.685757 1	2	10	1	1
	y ve	ector			

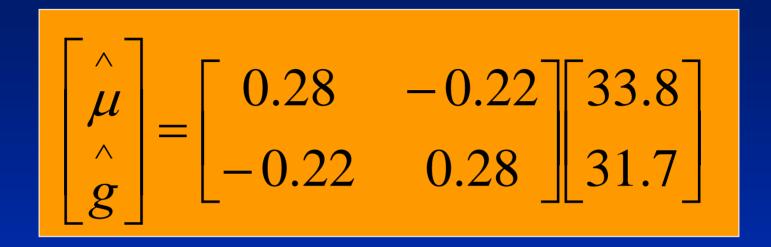
Estimate the marker effect and the mean as:



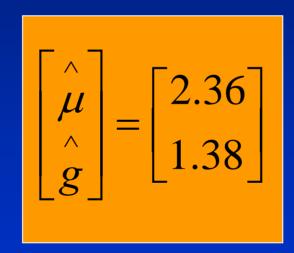




$$\begin{bmatrix} \land \\ \mu \\ \land \\ g \end{bmatrix} = \begin{bmatrix} 10 & 8 \\ 8 & 10 \end{bmatrix}^{-1} \begin{bmatrix} 33.8 \\ 31.7 \end{bmatrix}$$



• Estimates of the mean and marker effect are:

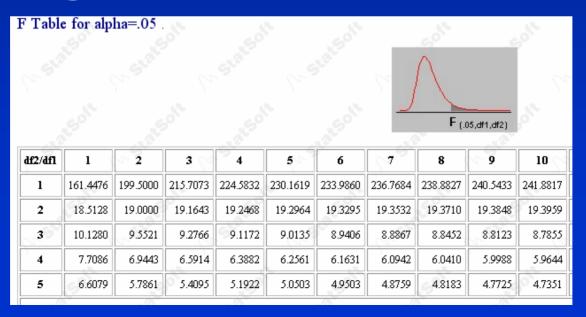


 In the "simulation", mean was 2, r² between QTL and marker was 1, and effect of 2 allele at QTL was 1.

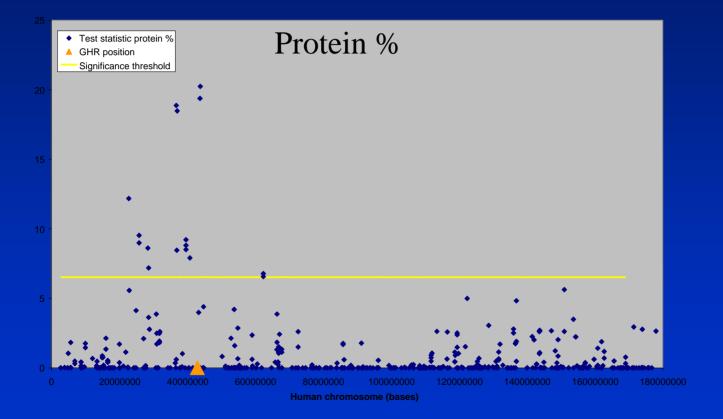
- Is the marker effect significant?
- F statistic comparing between marker variance to within marker variance
- Test against tabulated value for $F_{\alpha,v1,v2}$
 - $-\alpha =$ significance value
 - -v1=1 (1 marker effect for regression)
 - -v2=9 (number of records -1)

• In our simple example

 $-F_{data} = 4.56$ $-F_{0.05,1,9} = 5.12$ • Not significant



Results of genome scans with dense SNP panels





- 384 Holstein-Friesian dairy bulls selected from Australian dairy bull population
- genotyped for 10 000 SNPs
- Single marker regression with protein%

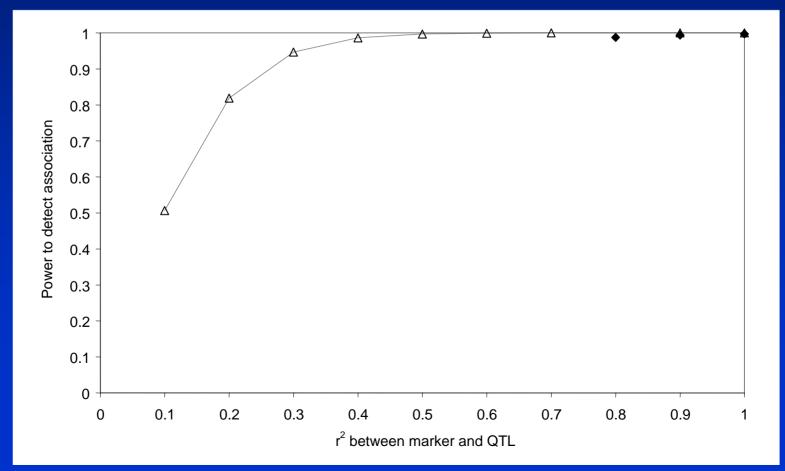


- What is the power of an association test with a certain number of records to detect a QTL?
- Power is probability of correctly rejecting null hypothesis when a QTL of really does exist in the population
- How many animals do we need to genotype and phenotype?

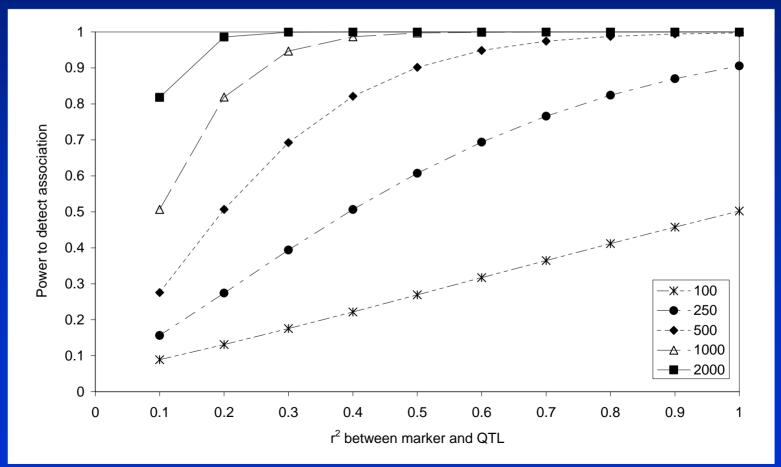
- Power is a function of:
 - r² between the marker and QTL
 - sample size must be increased by 1/r² to detect an un-genotyped QTL, compared with sample size for testing QTL itself
 - Proportion of total phenotypic variance explained by the QTL
 - Number of phenotypic records
 - Allele frequency of the rare allele of SNP
 - determines the minimum number of records used to estimate an allele effect.
 - The power becomes particular sensitive with very low frequencies (eg. <0.1).
 - The significance level α set by the experimenter

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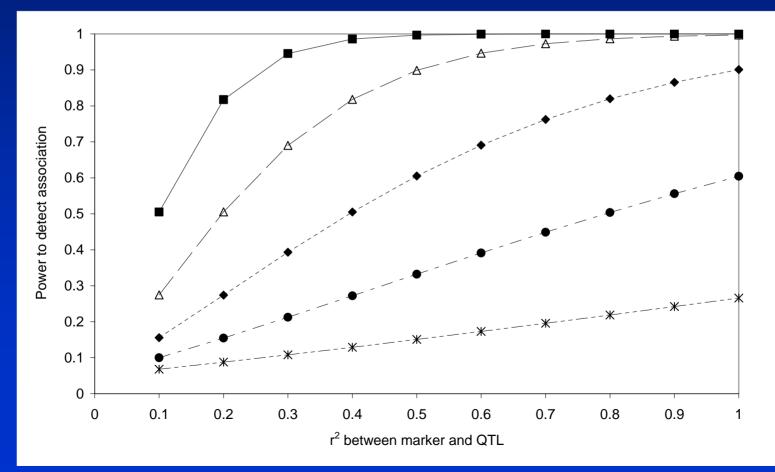
Power to detect a QTL explaining 5% of the phenotypic variance, 1000 phenotypic records



 Power to detect a QTL explaining 5% of the phenotypic variance

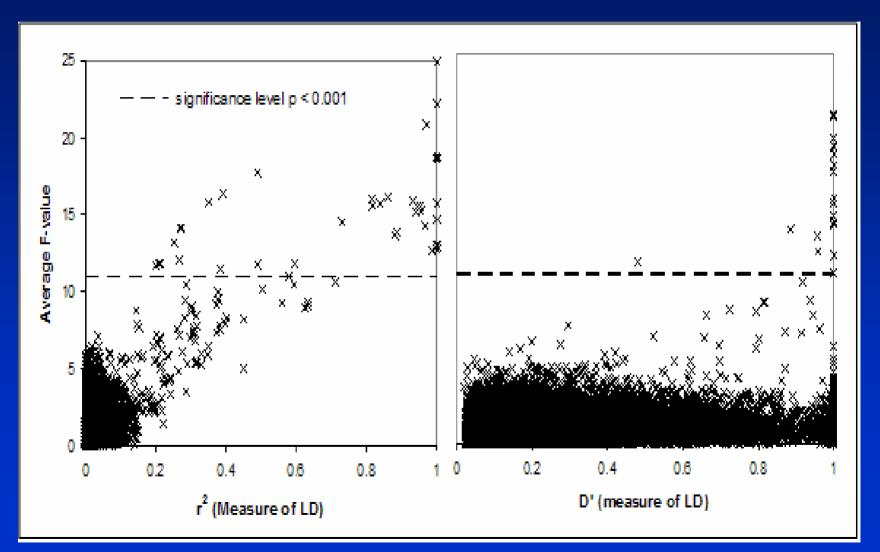


 Power to detect a QTL explaining 2.5% of the phenotypic variance



- r^2 of at least 0.2 is required to achieve power ≥ 0.8 to detect a QTL of $h_{QTL}=0.05$ with 1000 phenotypic records.
- In dairy cattle, $r^2 \approx 0.2$ at 100kb.
- Assuming a genome length of 3000Mb in cattle, we would need at least 15 000 markers to ensure there is a marker 100kb from every QTL.
- Assumes markers are evenly spaced, all have rare allele frequency > 0.2.
- In practise, markers not be evenly spaced, rare allele frequency of some markers below 0.2.
- At least 30 000 markers required.

- Another illustration of effect of r² on power
- An experiment to assess power of whole genome association scans in outbred livestock with commercially available SNP panels
 - 384 Angus cattle genotyped for 10,000 SNPs
 - QTL, polygenic and environmental effects were simulated for each animal
 - QTL simulated on genotyped SNPs chosen at random.
 - There was a strong correlation between F-value of significant SNPs and their r² with the "QTL"

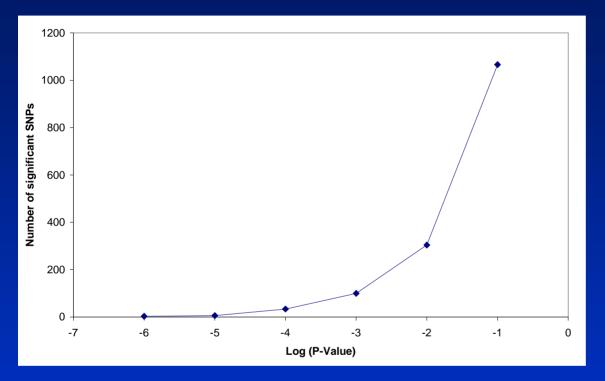


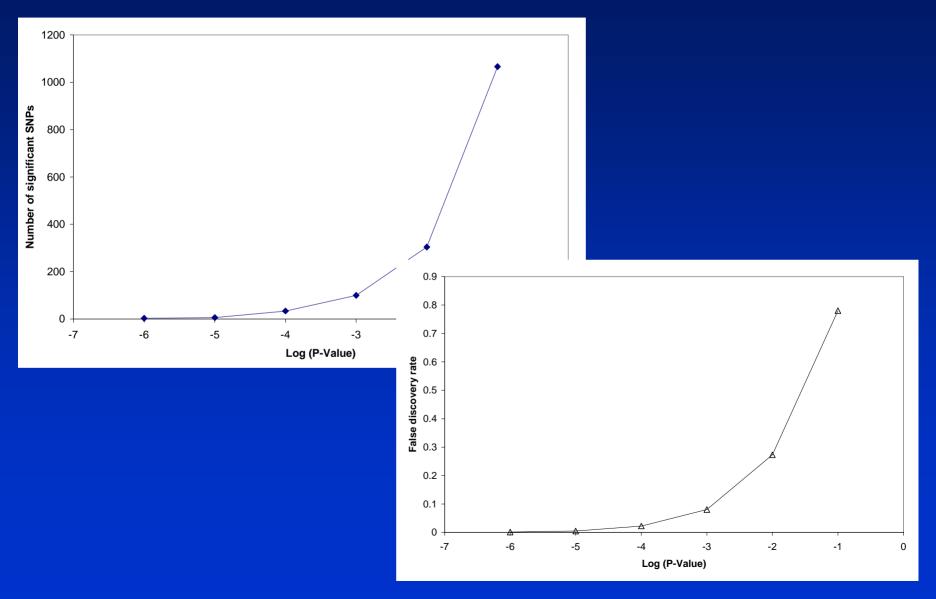
- What significance level to use?
 - P<0.01, P<0.001?
- We have a horrible multiple testing problem
 - Eg. If test 10 000 SNP at P<0.01 expect 100 significant results just by chance?
- Could just correct for the number of tests
 - But is too stringent, ignores the fact that tests are on the same chromosome (eg not independent)

- Could use a technique called permutation testing
 - Randomly shuffle phenotypes across genotypes
 - Test all SNPs (null hypothesis), get largest F value
 - Repeat 1000 times
 - 950th value is P<0.05 level corrected for multiple testing
- Difficult with pedigree structure

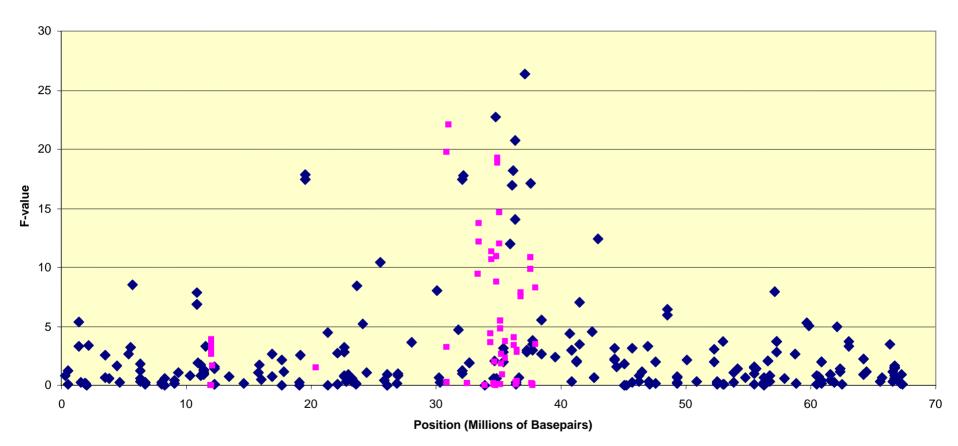
- An alternative is to choose a significance level with an acceptable false discovery rate (FDR)
- Proportion of significant results which are really false positives
- FDR = mP/n
 - m = number of markers tested
 - P = significance level (eg. P=0.01)
 - n = number of markers tested

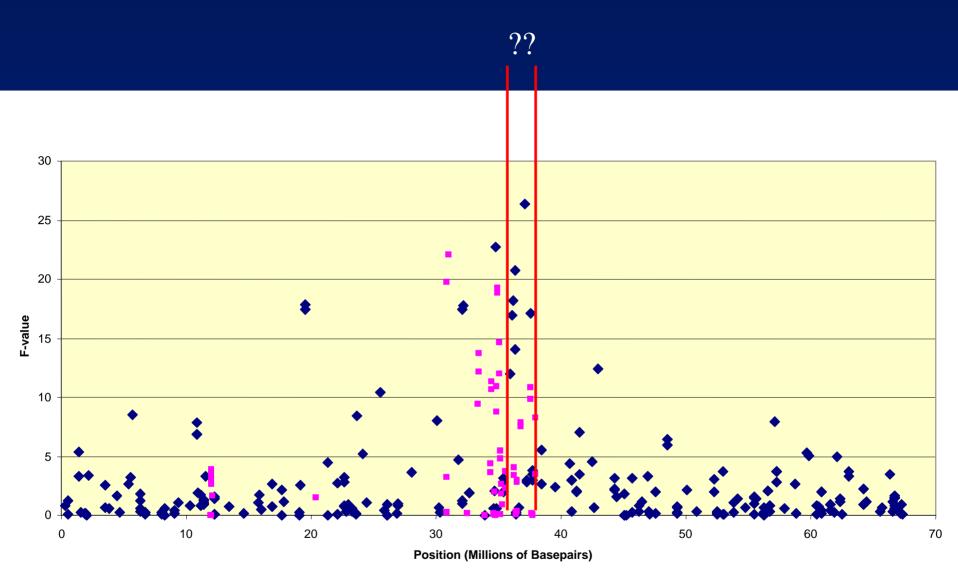
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 - n = number of markers tested
- Example
 - 10 000 markers tested at P<0.001, and 20 significant.
 What is FDR?
 - FDR=10000*0.001/20 = 50%
 - Eg. 50% of our significant results are actually false positives

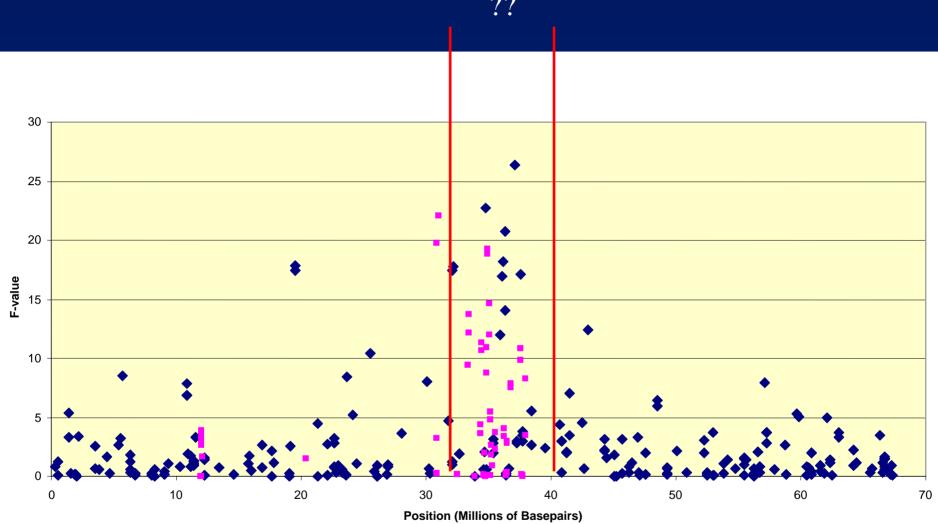




- Confidence regions
 - Following a genome wide association study, how do we decide the 95% confidence interval for the true QTL location?
- How many candidate genes to investigate?







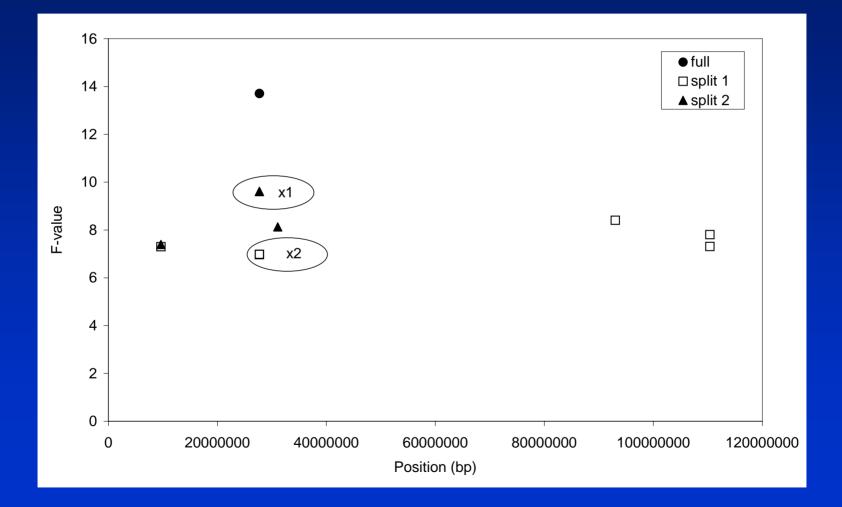
- One method to calculate confidence intervals
 - Count up number of "clusters"=n
 - Split data set into two at random (eg. half animals in one set, other half in other set)
 - Designate best SNP at a cluster location in data set 1 and data set 2 as x_{1i}, x_{2i} .
 - Estimate standard error of position over best SNP as:

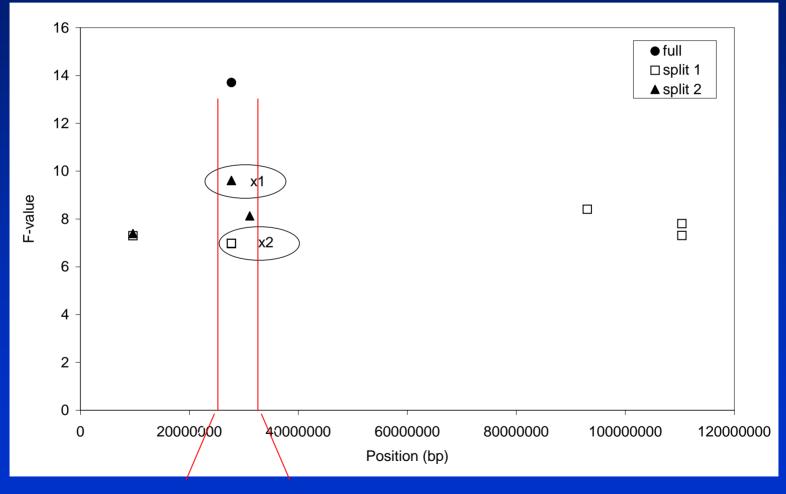
$$se(\bar{x}) = \sqrt{\frac{1}{4n} \sum_{i=1}^{n} (x_{1i} - x_{2i})^2}$$

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• 95% C.I = position of best SNP ±1.96*se(x_bar)





95% C.I

Mapping QTL using LD

- Association testing with single marker regression
- Accounting for population structure
- LD mapping with haplotypes
- The identical by descent (IBD) approach
- Combined linkage-linkage disequilibrium mapping

- Simple model we have used assumes all animals are equally (un) related.
- Unlikely to be the case.
- Multiple offspring per sire, breeds or strains all create population structure.
- If we don't account for this, false positives!

- Simple example
 - a sire has many progeny in the population.
 - the sire has a high estimated breeding value
 - a rare allele at a random marker is homozygous in the sire (*aa*)
 - Then sub-population of his progeny have higher frequency of a than the rest of the population.
 - As the sires' estimated breeding value is high, his progeny will also have higher than average estimated breeding values.
 - If we don't account for relationship between progeny and sire the rare allele will appear to have a (perhaps significant) positive effect.

• Can account for these relationships by extending our model.....

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}}' \boldsymbol{\mu} + \mathbf{X}g + \mathbf{Z}\mathbf{u} + \mathbf{e}$$

- Where
 - **u** is a vector of polygenic effect in the model with a covariance structure $u \sim N(0, A\sigma_a^2)$
 - A is the average relationship matrix built from the pedigree of the population
 - Z is a design matrix allocating animals to records.

• Can account for these relationships by extending our model.....

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}}' \boldsymbol{\mu} + \mathbf{X}g + \mathbf{Z}\mathbf{u} + \mathbf{e}$$

• Solutions ($\lambda = \sigma_e^2 / \sigma_a^2$):

$$\begin{bmatrix} \hat{\mu} \\ \hat{\mu} \\ \hat{g} \\ \hat{u} \\ u \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X & \mathbf{1}_{n}'Z \\ \mathbf{X'1}_{n} & \mathbf{X'X} & \mathbf{X'Z} \\ \mathbf{Z'1}_{n} & \mathbf{Z'X} & \mathbf{Z'Z} + \mathbf{A}^{-1}\lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{X'y} \\ \mathbf{Z'y} \end{bmatrix}$$

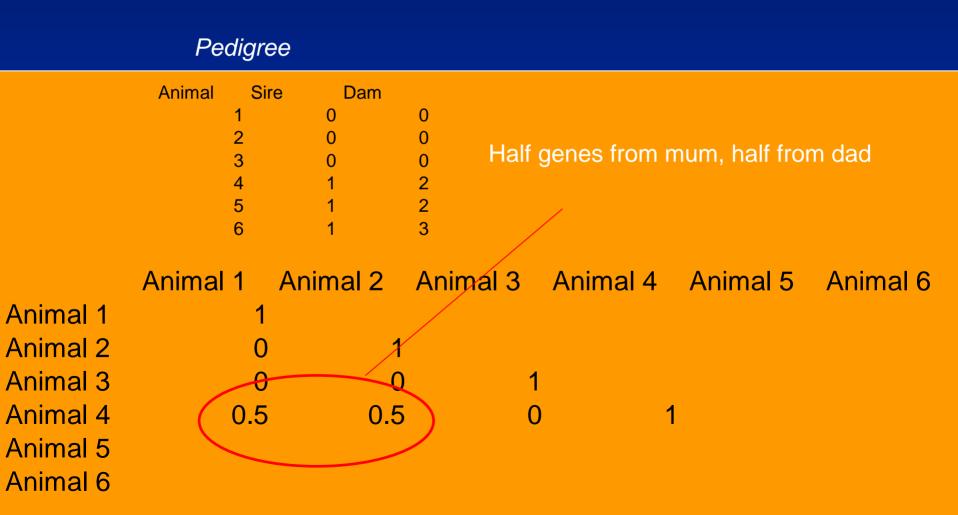
Pedigree

Animal	Sire	Dam	
	1	0	0
	2	0	0
	3	0	0
	4	1	2
	5	1	2
	6	1	3

	Pedig	ree				
	Animal 5 1 2 3 4 5 6	Sire Dam 0 0 1 1 1 1	0 0 0 2 2 3			
Animal 1 Animal 2 Animal 3 Animal 4 Animal 5 Animal 6		Animal 2 1	Animal 3	Animal 4	Animal 5	Animal 6

	Pedi	igree					
	Animal	Sire	Dam				
	1		0	0			
	2	2	0	0			
	3	3	0	0			
	4		1	2			
	5		1	2			
	6	5	1	3			
Animal 1 Animal 2 Animal 3 Animal 4 Animal 5 Animal 6	Animal 1	1 0	nimal 2 1	Animal 3	Animal 4	Animal 5	Animal 6

	iree				
Animal	Sire Dam				
1	0	0			
2	0	0			
		0			
		2			
		2			
6	1	3			
	1 0	Animal 1 0	3 Animal 1	4 Animal 5	Animal 6
	1 2 3 4 5 6 Animal 1	1 0 2 0 3 0 4 1 5 1 6 1	1 0 0 2 0 0 3 0 0 4 1 2 5 1 2 6 1 3 Animal 1 Animal 2 Animal 1 0 1	1 0 0 2 0 0 3 0 0 4 1 2 5 1 2 6 1 3 Animal 1 Animal 2 Animal 3 Animal 1 0 1	1 0 0 2 0 0 3 0 0 4 1 2 5 1 2 6 1 3 Animal 1 Animal 2 Animal 3 Animal 4 Animal 5 1 0 1



Dodiaroo

	Peo	aigre	e					
	Animal	Sir	e Dam					
		1	0	0				
		2	0	0				
		3	0	0				
		4	1	2				
		5	1	2				
		6	1	3				
	Animal	1	Animal 2	Animal 3		Animal 4	Animal 5	Animal 6
Animal 1		1						
Animal 2		0	1					
Animal 3		0	0		1			
Animal 4		0.5	0.5		C	1		
Animal 5		0.5	0.5		0	0.5	1	
Animal 6								

	Ped	ligree	9								
	Animal	Sire	Dam								
		1	0	0							
		2	0	0							
		3	0	0	Anim	nals 4	and 5	are ful	I SIDS		
		4	1	2							
		5	1	2							
		6	1	3							
	Animal	1 /	Animal 2	An	imal 3	Ani	mal 4	Anii	mal 5	Animal	6
Animal 1		1									
Animal 2		0	1								
Animal 3		0	0		1						
Animal 4		0.5	0.5		0)		4			
Animal 5		0.5	0.5		0)	(0.	5)	1	l	
Animal 6											

	Pedig	ree							
	Animal S	Sire I	Dam						
	1	0	0						
	2	0	0	A i		- 16 - 11		-	
	3	0	0	Anim	ials 6 is a h	ait si	o of 4 and	5	
	4	1	2						
	5	1	2						
	6	1	3						
	Animal 1	Anima	al 2 An	imal 3	Animal 4	A	nimal 5	Animal	6
Animal 1	· · · · · · · · · · · · · · · · · · ·	1							
Animal 2	(C	1						
Animal 3	(C	0	1					
Animal 4	0.8	5	0.5	0)	1			
Animal 5	0.8	5	0.5	0	0	.5	1		
Animal 6	0.8	5	0	0.5	0.2	25	0.25	>	-

Animal	Sire	Dam	Ph	nenotype SN	P allele SNI	P allele
	1	0	0	10.1	1	2
	2	0	0	2.2	2	2
	3	0	0	2.31	2	2
	4	1	2	6.57	1	2
	5	1	2	6.06	1	2
	6	1	3	6.21	1	2

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}} \boldsymbol{\mu} + \mathbf{X} g + \mathbf{e}$$

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$$\begin{bmatrix} \land \\ \mu \\ \land \\ g \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X \\ \mathbf{X'1}_{n} & \mathbf{X'X} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{X'y} \end{bmatrix}$$

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$$\mathbf{y} = \mathbf{1}_{\mathbf{n}} \boldsymbol{\mu} + \mathbf{X} g + \mathbf{e}$$

$$\begin{bmatrix} \land \\ \mu \\ \land \\ g \end{bmatrix} = \begin{bmatrix} 6 & 8 \\ 8 & 12 \end{bmatrix}^{-1} \begin{bmatrix} 33.5 \\ 38 \end{bmatrix}$$

Animal	Sire	Dam	Р	henotype SNI	P allele SNI	P allele
	1	0	0	10.1	1	2
	2	0	0	2.2	2	2
	3	0	0	2.31	2	2
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12.2 μ 5

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Animal	Sire	Dam	Ph	nenotype SNI	P allele SNF	^{>} allele
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$$\begin{bmatrix} \land \\ \mu \\ \land \\ \mathbf{g} \\ \land \\ \mathbf{u} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X & \mathbf{1}_{n}'Z \\ \mathbf{X'1}_{n} & \mathbf{X'X} & \mathbf{X'Z} \\ \mathbf{Z'1}_{n} & \mathbf{Z'X} & \mathbf{Z'Z} + \mathbf{A}^{-1}\lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{X'y} \\ \mathbf{Z'y} \end{bmatrix}$$

• Example

Animal	Sire	Dam	Pł	nenotype SN	P allele SNI	P allele
	1	0	0	6.51	1	1
	2	0	0	2.2	2	2
	3	0	0	2.31	2	2
	4	1	2	4.72	1	2
	5	1	2	5.02	1	2
	6	3	2	2.93	2	2

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}}'\boldsymbol{\mu} + \mathbf{X}g + \mathbf{Z}\mathbf{u} + \mathbf{e}$$

 $\lambda = 0.33$

Animal	Sire	Dam	Pł	nenotype SN	P allele SN	^{>} allele
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	4	1	2	4.72	1	2
	5	1	2	5.02	1	2
	6	3	2	2.93	2	2

		6	8	1	1	1	1	1	1	_1	33.45
U		8	12	1	2	2	1	1	1	–	37.96
		1	1	1.916575	0.416625	0.16665	-0.3333	-0.49995	-0.3333		10.1
σ	=	1	2	0.416625	1.749925	0	-0.3333	-0.49995	0		2.2
		1	2	0.16665	0	1.49995	0	0	-0.3333		2.31
		1	1	-0.49995	-0.49995	0	1.6666	0.3333	0		6.57
		1	1	-0.3333	-0.3333	0	0	1.6666	0		6.06
<u> </u>	J	1	1	-0.3333	0	-0.3333	0	0	1.6666		6.21

Animal	Sire	Dam	Pr	nenotype SNI	P allele SNI	P allele
	1	0	0	10.1	1	2
	2	0	0	2.2	2	2
	3	0	0	2.31	2	2
	4	1	2	6.57	1	2
	5	1	2	6.06	1	2
	6	1	3	6.21	1	2

$$\begin{bmatrix} 10.3 \\ -3.5 \\ 1.9 \\ -1.1 \\ -0.9 \\ 0.2 \\ -0.3 \\ -0.1 \end{bmatrix}$$

- Example of importance of accounting for population structure.....
 - 365 Angus cattle genotyped for 10,000 SNPs
 - polygenic and environmental effects were simulated for each animal
 - No QTL fitted!
 - Effect of each SNP tested using three models
 - SNP only
 - SNP and sire
 - SNP and full pedigree

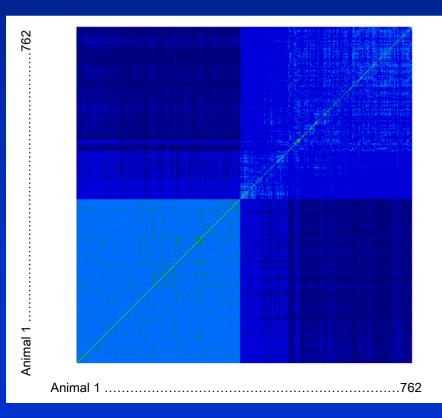
Number of false positives......

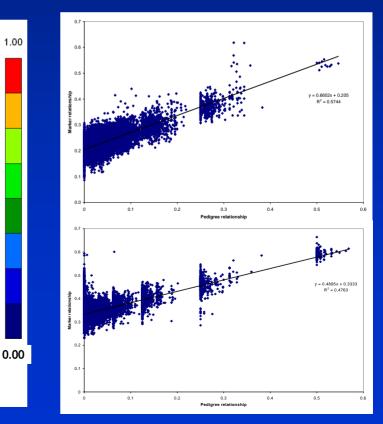
Analysis model	Significance level				
	p<0.005	p<0.001	p<0.0005		
Expected type I errors	40	8	4		
1. Full pedigree model	39 (SD=14)	9 (SD=5)	4 (SD=3)		
2. Sire pedigree model	46 [*] (SD=21)	11 [*] (SD=7)	6 [*] (SD=5.5)		
3. No pedigree model	68 ^{**} (SD=31)	18 ^{**} (SD=11)	10 ^{**} (SD=7)		
4. Selected 27% - full pedigree	54 ^{**} (SD=18)	12 ^{**} (SD=6)	7 ^{**} (SD=4)		

Problem when we do not have history of the population

1.00

Solution – use the average relationship across all markers as the **A** matrix





Mapping QTL using LD

- Association testing with single marker regression
- Accounting for population structure
- LD mapping with haplotypes
- The identical by descent (IBD) approach
- Combined linkage-linkage disequilibrium mapping

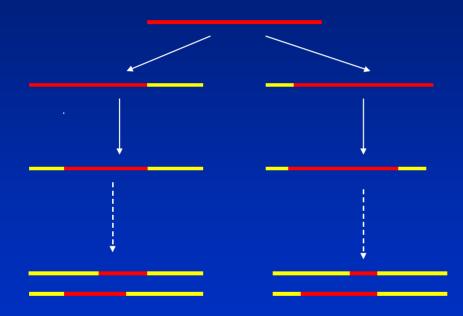
LD mapping with haplotypes

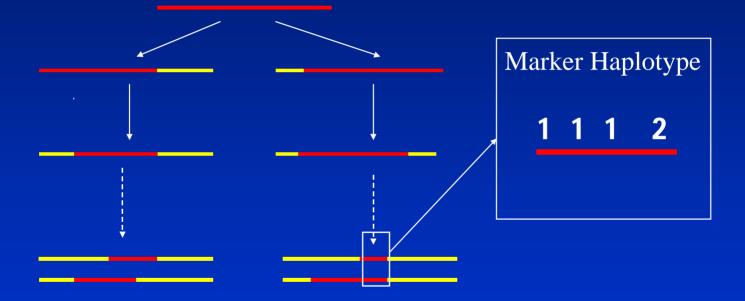
- Power of association study depends on LD between markers and QTL
- One way to increase LD between QTL alleles and markers is to use *haplotypes* of markers rather than a single marker
- 1_Q single marker (1 is the allele of the marker)
- 1_1_Q_2_1 Haplotype of markers

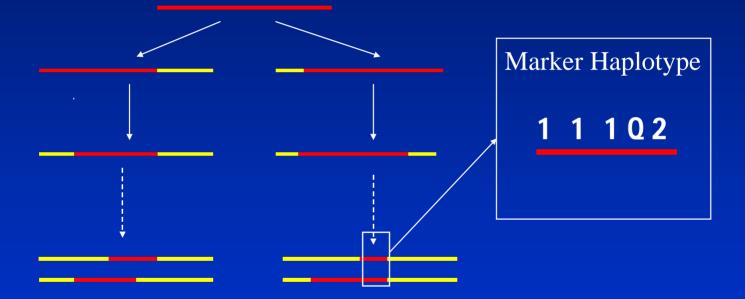
LD mapping with haplotypes

- Value of haplotypes depends on LD between haplotype and QTL
 - If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
 - If probability is high, high level of LD between haplotype and QTL

- If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
- Haplotypes identical either because chromosome segments from same common ancestor





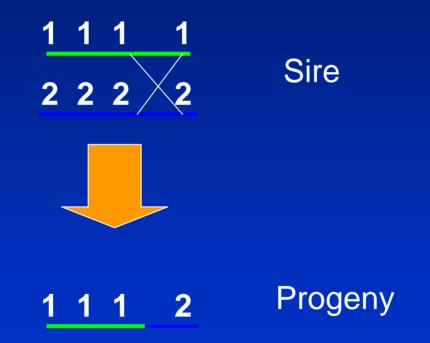


- If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
- Haplotypes identical either because chromosome segments from same common ancestor
- Or because of chance recombination.....



Sire

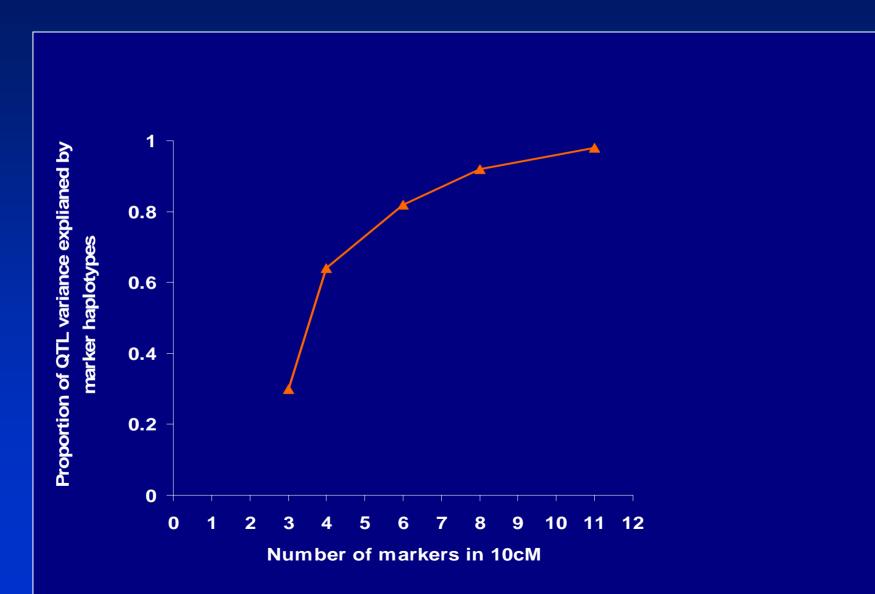
Formation of gamete



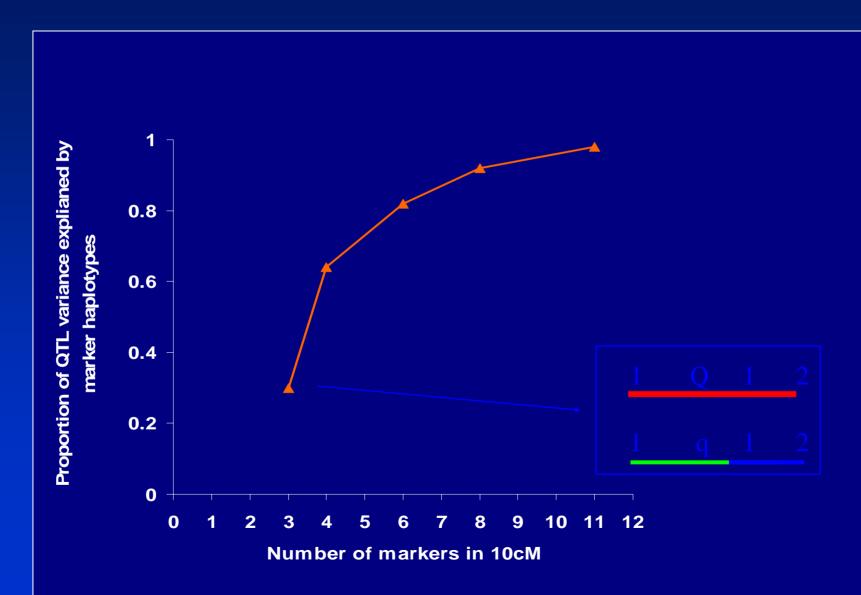




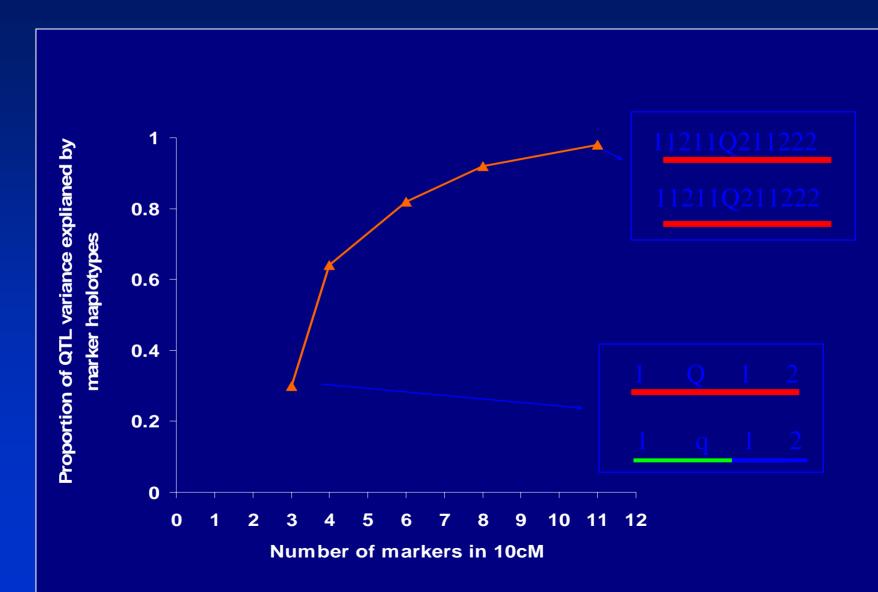
Proportion of QTL variance explained by surrounding markers



Proportion of QTL variance explained by surrounding markers



Proportion of QTL variance explained by surrounding markers



- If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
- Haplotypes identical either because chromosome segments from same common ancestor
- Or because of chance recombination......
- With more markers in haplotype, the chance of creating the same haplotype by recombination becomes small

• Model ?

$$y = 1_n' \mu + Xg + Zu + e$$

- Where g is now a vector of haplotype effects dimensions (number of haplotypes observed x 1)
- And X allocates records to haplotyes

Example (eg after using PHASE to infer haplotype)

Animal	Paternal haplotype	Maternal haplotype	
	1	1	1
	2	1	2
:	3	2	3
	4	5	4
	5	3	2



Example (eg after using PHASE to infer haplotype)

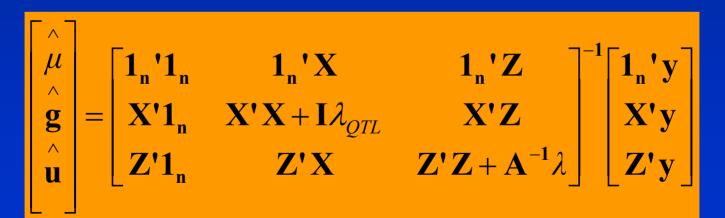
	Animal	Paternal hap	lotype	Maternal	haplotype	Э	
	1		1			1	
	2		1			2	
	3		2			3	
	4		5			4	
	5		3			2	
			Haplotype				
• X		1		2	3	4	5
	1	2		0	0	0	0
	2	2 1		1	0	0	0
Animal	3	0		1	1	0	0
	۷	0		0	0	1	1
	5	5 0		1	1	0	0

A

- Fit haplotypes as random effects
 - **g** ~ N(0, $\sigma_{\rm h}^2$)
 - Some haplotypes will be rare, very few observations
 - Fitting the haplotype effect as random regresses the effects back to account for the lack of information

$$\begin{bmatrix} \land \\ \mu \\ \land \\ \mathbf{g} \\ \land \\ \mathbf{u} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X & \mathbf{1}_{n}'Z \\ \mathbf{X'1}_{n} & \mathbf{X'X} & \mathbf{X'Z} \\ \mathbf{Z'1}_{n} & \mathbf{Z'X} & \mathbf{Z'Z} + \mathbf{A}^{-1}\lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{X'y} \\ \mathbf{Z'y} \end{bmatrix}$$

- Fit haplotypes as random effects
 - $g \sim N(0, \sigma_h^2)$
 - Some haplotypes will be rare, very few observations
 - Fitting the haplotype effect as random regresses the effects back to account for the lack of information
 - $\lambda_h = \sigma_e^2 / \sigma_h^2$



- There is a "cost" of using haplotypes instead of single markers
- With single markers only one effect to estimate, with haplotypes many effects
- Fewer observations per effect, lower accuracy of estimating each effect

	Proportion of	Maximum	Observed
	QTL variance	number of	number of
	explained	haplotypes	haplotypes
Nearest marker	0.10	2	2
Best marker	0.20	2	2
2 Marker haplotypes	0.15	4	3.4
4 Marker haplotypes	0.28	16	9.4
6 Marker haplotypes	0.55	64	20.8

Mapping QTL using LD

- Association testing with single marker regression
- Accounting for population structure
- LD mapping with haplotypes
- The identical by descent (IBD) approach
- Combined linkage-linkage disequilibrium mapping

• Principle:

 Existence of LD implies small segments of chromosome in population which are descended from the same common ancestor (IBD).

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- IBD chromosome segments will not only carry identical marker haplotypes; if there is a QTL within chromosome segment, IBD chromosome segments will also carry identical QTL alleles.

• Principle:

- Existence of LD implies small segments of chromosome in population which are descended from the same common ancestor (IBD).
- IBD chromosome segments will not only carry identical marker haplotypes; if there is a QTL within chromosome segment, IBD chromosome segments will also carry identical QTL alleles.
- If two animals carry chromosomes which are IBD at a QTL position, their phenotypes will be correlated.

• The model

$$y_i = \mu + u_i + vp_i + vm_i + e_i$$

- Where vp_i and vm_i are the effects of the paternal and maternal QTL alleles respectively
- modelling the effect of the QTL directly rather than assuming a haplotype or marker is in LD with the QTL

- Each animal has it's own QTL alleles
- There is a probability that different QTL alleles are actually IBD
- This is captured in the IBD (G) matrix
- Elements g_{ij} is the probability that QTL allele i and j are IBD.
- This probability is inferred from marker haplotypes
- Dimensions (2*number of animals * 2*number of animals)
- $u \sim (0, \mathbf{A}\sigma_a^2), v \sim (0, \mathbf{G}\sigma_v^2), e \sim \sim (0, \|\sigma_e^2)$

- Building IBD matrix from marker haplotypes
 - Consider three haplotypes drawn from population at random (P is putative QTL position)
 - A 112P112
 - B 2<u>12P112</u>
 - C 222P222
 - P(IBD at QTL A,B) >P(IBD at QTL B,C), as longer identical haplotype

- Building IBD matrix from marker haplotypes
 - Parameters which determine IBD coefficients are
 - extent of LD
 - length of haplotype and
 - number of markers in the haplotype

- Building IBD matrix from marker haplotypes
- Algorithm of Meuwissen and Goddard (2001)
 - deterministically predicts IBD coefficients at putative QTL positions from marker haplotypes

- Building IBD matrix from marker haplotypes
- Algorithm of Meuwissen and Goddard (2001)
 - deterministically predicts IBD coefficients between two marker haplotypes using
 - number of markers flanking QTL position which are identical by state
 - probability identical by chance ~ marker homozygosity
 - extent of LD based on length of haplotype, effective population size

- Building IBD matrix from marker haplotypes
 - An example with Ne = 100
 - 6 markers in 10cM, putative QTL position in centre M_M_M_Q_M_M_M
 - Sample four haplotypes from the population
 - 112112, 112112, 122112, 222122
 - IBD matrix is:

	112112	112112	122112	222122
112112	1			
112112	0.82	1		
122112	0.63	0.63	1	
222122	0.49	0.49	0.56	1

- A two stage approach for linkage disequilibrium mapping
 - For each putative QTL position, IBD or G matrix. IBD matrix has elements g_{ij}=Prob(QTL alleles i and j are identical by descent or IBD)
 - For each position considered in step 1, construct the linear model to estimate QTL variances and other parameters, test for presence of QTL

• The model

$$y_i = \mu + u_i + vp_i + vm_i + e_i$$

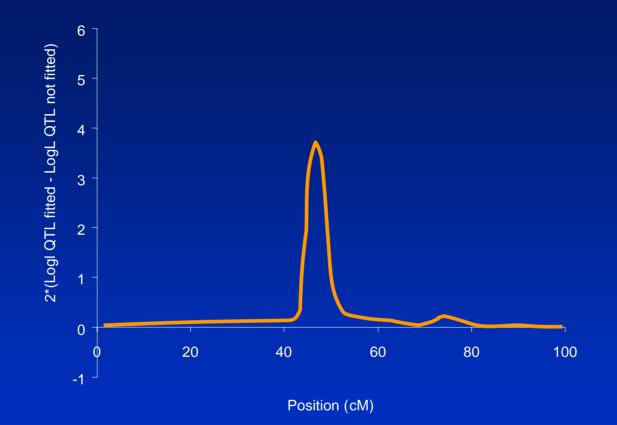
• $u \sim (0, A\sigma_a^2), v \sim (0, G\sigma_v^2), e \sim (0, I\sigma_e^2)$

- Use variance component estimation procedures to find the
 - Estimate of $\sigma_{\rm u}{}^2$
 - Estimate of $\sigma_{\rm v}{}^2$
 - Estimate of $\sigma_{e}{}^{2}$
 - Which maximise the Log likelihood (LogL) of the data given these parameters
 - Eg. ASREML

- How do we test if the QTL is significant or not?
- Fit the model with no QTL:

$$y_i = \mu + u_i + e_i$$

 Plot -2*(LogL QTL fitted - LogL QTL not fitted) against position



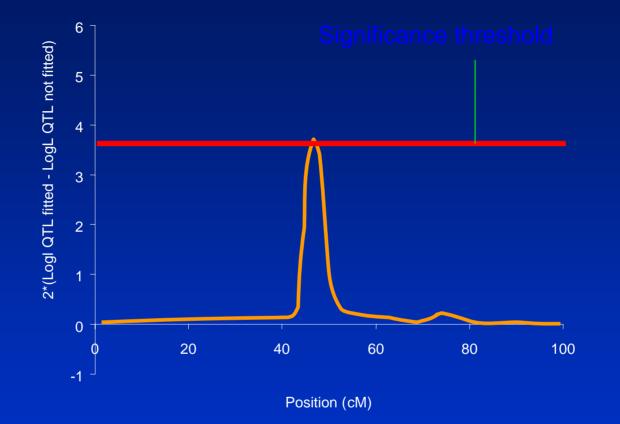
The IBD approach

- How do we test if the QTL is significant or not?
- Fit the model with no QTL:

$$y_i = \mu + u_i + e_i$$

- Plot -2*(LogL QTL fitted LogL QTL not fitted) against position
- Is distributed as a $\chi^2_{1, \alpha}$ where α is the desired significance level
- at α=0.05 is 3.84)

Linkage mapping in complex pedigrees



The IBD approach

Confidence interval Drop of 2 of test statistic from maximum point

Comparison of approaches

- Zhao et al. (2007) compared power and precision of QTL mapping with single marker regression, haplotypes and IBD approach
- They found in simulated data that single marker regression and the IBD approach had similar power and precision
- Calus et al. (2007) found that haplotypes gave slightly greater accuracy than single markers, and that the IBD approach gave much higher accuracies at low marker densities
- Hayes et al. (2007) tried to use real data, and results indicated 6 marker haplotypes were better than single marker regression
- Level of LD, simulation assumptions??

Mapping QTL using LD

- Association testing with single marker regression
- Accounting for population structure
- LD mapping with haplotypes
- The Identical by descent (IBD) approach
- Combined linkage-linkage disequilibrium mapping

- Extent of LD very variable
- LD can exist between loci on different chromosomes!!
- Combine LD and linkage information to filter spurious peaks

- Consider a half sib design
 - LD information from sire haplotypes, maternal hapotypes of progeny
 - Linkage information from paternal haplotypes of progeny
- IBD matrix:

	SH	MHP	PHP
SH	[a]	[a]	[b]
MHP	[a]	[a]	[b]
PHP	[b]	[b]	[b]

- a = LD (Meuwissen and Goddard 2001)

-b = linkage

• In linkage analysis (LA) consider founder alleles (sires, dams) to be unrelated, eg.....

Sire 1211, 1212 Dam 1222, 1211 Progeny 1211, 1222 LA

		Sire		Dam		Progeny	
		Pat	Mat	Pat	Mat	Pat	Mat
Sire	Pat	1					
	Mat	0	1				
Dam	Pat	0	0	1			
	Mat	0	0	0	1		
Progeny	Pat	1	0	0	0	1	
	Mat	0	0	1	0	0	1

Sire 1211, 1212 Dam 1222, 1211 Progeny 1211, 1222 LA

		Sire		Dam		Progeny	
		Pat	Mat	Pat	Mat	Pat	Mat
Sire	Pat	1					
	Mat	0	1				
Dam	Pat	0	0	1			
	Mat	0	0	0	1		
Progeny	Pat	1	0	0	0	1	
	Mat	0	0	1	0	0	1

LD

		Sire		Dam	
		Pat	Mat	Pat	Mat
Sire	Pat	1			
	Mat	0.8	1		
Dam	Pat	0.5	0.5	1	
	Mat	0.9	0.5	0.5	1

Sire 1211, 1212 Dam 1222, 1211 Progeny 1211, 1222 LA

		Sire		Dam		Progeny	
		Pat	Mat	Pat	Mat	Pat	Mat
Sire	Pat	1					
	Mat	0	1				
Dam	Pat	0	0	1			
	Mat	0	0	0	1		
Progeny	Pat	1	0	0	0	1	
	Mat	0	0	1	0	0	1

LD

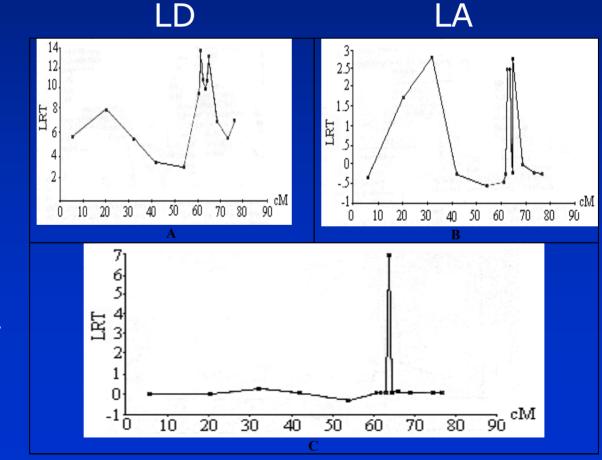
		Sire		Da	am
		Pat	Mat	Pat	Mat
Sire	Pat	1			
	Mat	0.8	1		
Dam	Pat	0.5	0.5	1	
	Mat	0.9	0.5	0.5	1

Sire 1211, 1212 Dam 1222, 1211 Progeny 1211, 1222

LDLA

		Sire		Dam		Progeny	
		Pat	Mat	Pat	Mat	Pat	Mat
Sire	Pat	1					
	Mat	0.8	1				
Dam	Pat	0.5	0.5	1			
	Mat	0.9	0.5	0.5	1		
Progeny	Pat	1	0.8	0.5	0.9	1	
	Mat	0.5	0.5	1	0.5	0.8	1

Combined LD-LA mapping
Example of twinning QTL in Norwegian dairy cattle (Meuwissen et al. 2002)

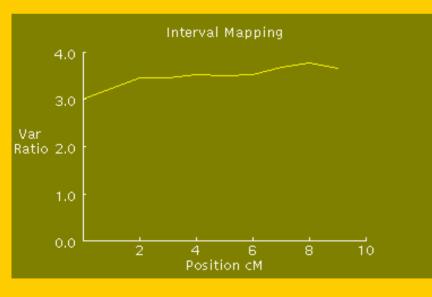


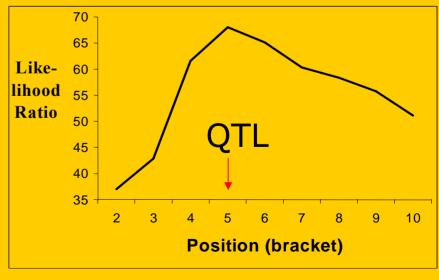
LD-LA

- How much information does LD add to the analysis?
 - Depends on marker spacing and extent of LD

LA

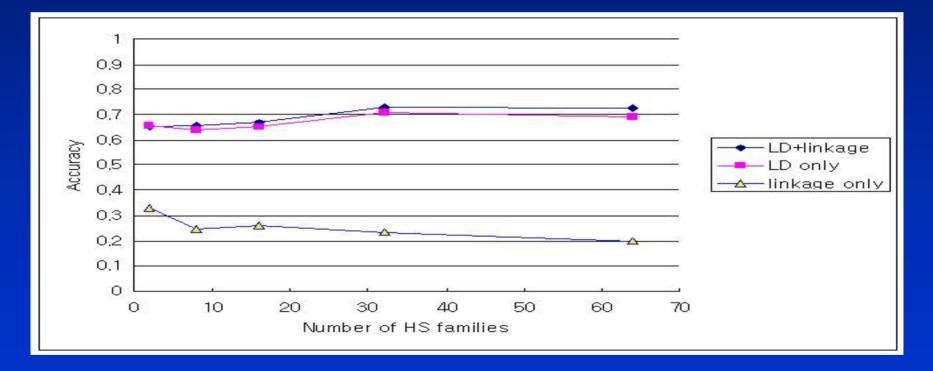
LD-LA





A

Combined LD-LA mapping Can we use half-sib families for LD analysis?



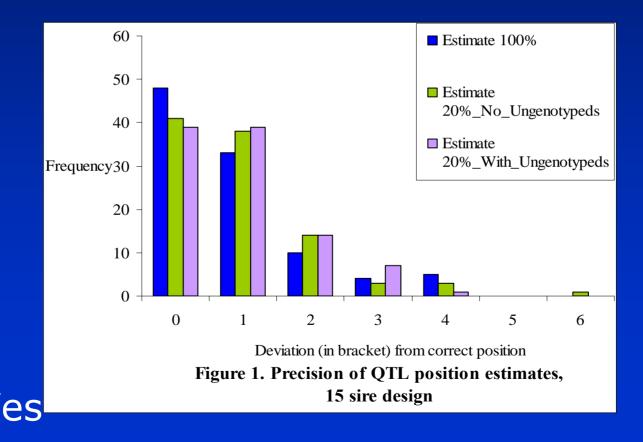
Yes
 – Dam haplotypes provide LD information

 Number of progeny required to position QTL to a 95% C.I. 3cM interval with different designs:

Population	Number of genotyped progeny required to map QTL to 3cM 95% C.I.
F2	7407
Full sib	12685
Commercial (LDLA)	900

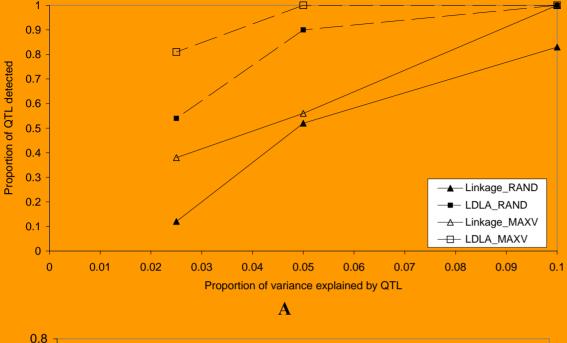
• Of course depends on assumptions about extent of LD $\,\sim\,$ determined by $\rm N_e$

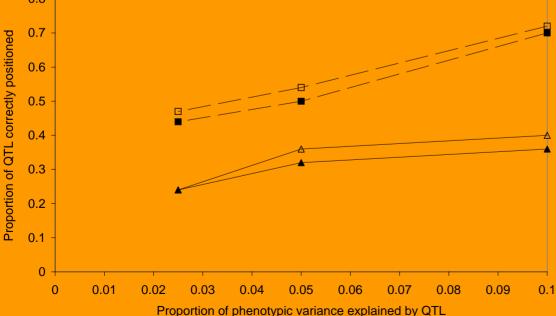
- Can we use half-sib families for LD analysis?
 - + selective genotyping ?



 LDLA analysis + selective genotyping = Cheap? experiment able to position QTL with high degree of precision

- Which families to pick for LD analysis?
 - Those with maximum within half sib family variance
 - Maximise chance of QTL segregating
 - Another `selective genotyping' strategy





Power of a mapping design (30 sires mated to 60 dams in a half sib design and 10 progeny per dam, no recombination in males) to

A. detect QTL andB. Position QTL within 3cM of the true QTL position,

for QTL explaining different proportions of the phenotypic variance. The 30 half sib families were either randomly selected (RAND) from the breeding population for genotyping, or the half sib families with the largest within half sib family variance for the trait (MAXV) were selected.

LD mapping of QTL

Take home points

- LD mapping uses information on historical recombinants to narrow QTL C.I.
- Power depends on extent of LD and marker density
 - Knowledge of extent of LD critical
- Some suggestion that single marker regression a good approach, with high marker density?
- IBD approach allows extension to capture LA information
 - v. important with lower marker density >> power
 - filter spurious peaks
- Half sib designs ideal for LDLA mapping
 - Use LD info from dam haplotypes